

# Structural Estimation of Time-varying Spillovers: an Application to International Credit Risk Transmission

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# Introduction – Motivation

- Numerous econometric models to gauge financial spillovers (Dungey *et al.*, 2005).

**Main issue: endogeneity of asset prices.**

- Among them, **Diebold and Yilmaz** (2009, DY) allow for the construction of time-varying spillovers, from pairwise to systemwise.
- Yet, DY-approach relies on identified SVAR, and empirical research:
  - Either **circumvents** identification-issues (Cholesky, GIRF, Claeys and Vašíček, 2014).
  - Or has attractive identification, but **no time variation** of spillovers **outside of rolling window estimation** (Ando *et al.*, 2018, De Santis and Zimic, 2018, 2019).

# Introduction – Overview of our approach

## General idea:

- Exploit time-varying FEVDs as measures of spillovers (Diebold-Yilmaz).

## Data:

- 9 EZ sovereign CDS: BE DE ES FR GR IE IT NL PT.
- 7 corresponding EZ bank CDS, apart for GR and IE. Daily since 2008.

## Econometric roadmap:

- 1 Filter CDS series from **common shocks**.
- 2 Estimate a SVAR-GARCH on the residuals, **identified by heteroskedasticity**.
- 3 **Economic identification** by contribution, validated by matching structural shocks with historical events.
- 4 TV variances from GARCH processes allow construction of TV FEVDs.

## ① Methodological Contribution

- Identified orthogonal shocks even in a **16-variable** SVAR.
- Our contagion indices appear **more reactive** than other methods used in the literature (Granger-causality).
- Pairwise spillovers **better identify** the sources of shocks.

## ② Economic Validation

- Our estimates match both **the narratives of spillovers** during the EZ debt crisis...
- ...and **the theoretical channels** of credit risk spillovers.

# Model – Identification by Heteroskedasticity

$$\begin{array}{ll} \text{Reduced form VAR} & \text{Structural VAR} \\ \mathbf{y}_t = \mathbf{A}(\mathbf{L}) \mathbf{y}_{t-1} + \boldsymbol{\mu}_t & \rightarrow \mathbf{B}_0 \mathbf{y}_t = \mathbf{B}(\mathbf{L}) \mathbf{y}_{t-1} + \boldsymbol{\epsilon}_t \end{array}$$

Assume:

$$E(\boldsymbol{\mu}_t \boldsymbol{\mu}_t') = \begin{cases} \boldsymbol{\Sigma}_1 & \text{if } t \leq T \\ \boldsymbol{\Sigma}_2 & \text{if } t > T \end{cases} \quad E(\boldsymbol{\epsilon}_t \boldsymbol{\epsilon}_t') = \begin{cases} \mathbf{I} & \text{if } t \leq T \\ \boldsymbol{\lambda} & \text{if } t > T \end{cases}$$

Then, fully identified system:

$$\boldsymbol{\Sigma}_1 = \mathbf{B}_0 \mathbf{B}_0^{-1} \quad \text{and} \quad \boldsymbol{\Sigma}_2 = \mathbf{B}_0 \boldsymbol{\lambda} \mathbf{B}_0^{-1}$$

Works also with more than two volatility regimes (as in our case)

# Model – SVAR-GARCH

Assume:

$$\epsilon_{k,t} = \sigma_{k,t|t-1} e_{k,t} \quad \text{where } \mathbf{e}_t \sim \text{i.i.d. } N(\mathbf{0}, \mathbf{I}_N) \quad \text{and} \quad (1)$$

$$\sigma_{k,t|t-1}^2 = (1 - \gamma_k - g_k) + \gamma_k (\epsilon_{k,t-1})^2 + g_k \sigma_{k,t-1|t-2}^2 \quad (2)$$

# Model – SVAR-GARCH

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$$\epsilon_{k,t} = \sigma_{k,t|t-1} e_{k,t} \quad \text{where } \mathbf{e}_t \sim \text{i.i.d. } N(\mathbf{0}, \mathbf{I}_N) \quad \text{and} \quad (1)$$

$$\sigma_{k,t|t-1}^2 = (1 - \gamma_k - g_k) + \gamma_k (\epsilon_{k,t-1})^2 + g_k \sigma_{k,t-1|t-2}^2 \quad (2)$$

Then, we can express the reduced form shocks as:

$$\boldsymbol{\mu}_t = \mathbf{B}_0^{-1} \boldsymbol{\lambda}_{t|t-1}^{\frac{1}{2}} \mathbf{e}_t \quad (3)$$

where:

$$\boldsymbol{\lambda}_{t|t-1} = \begin{bmatrix} \sigma_{1,t|t-1}^2 & & 0 \\ & \dots & \\ 0 & & \sigma_{N,t|t-1}^2 \end{bmatrix} \quad (4)$$

# Model – SVAR-GARCH

Assume:

$$\epsilon_{k,t} = \sigma_{k,t|t-1} e_{k,t} \quad \text{where } \mathbf{e}_t \sim \text{i.i.d. } N(\mathbf{0}, \mathbf{I}_N) \quad \text{and} \quad (1)$$

$$\sigma_{k,t|t-1}^2 = (1 - \gamma_k - g_k) + \gamma_k (\epsilon_{k,t-1})^2 + g_k \sigma_{k,t-1|t-2}^2 \quad (2)$$

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Statistical identification:

$$\boldsymbol{\Sigma}_{u,t|t-1} = \mathbf{B}_0^{-1} \boldsymbol{\lambda}_{t|t-1} \mathbf{B}_0^{-1'} \quad (5)$$

# Model – Diebold-Yilmaz Index (sum-up)

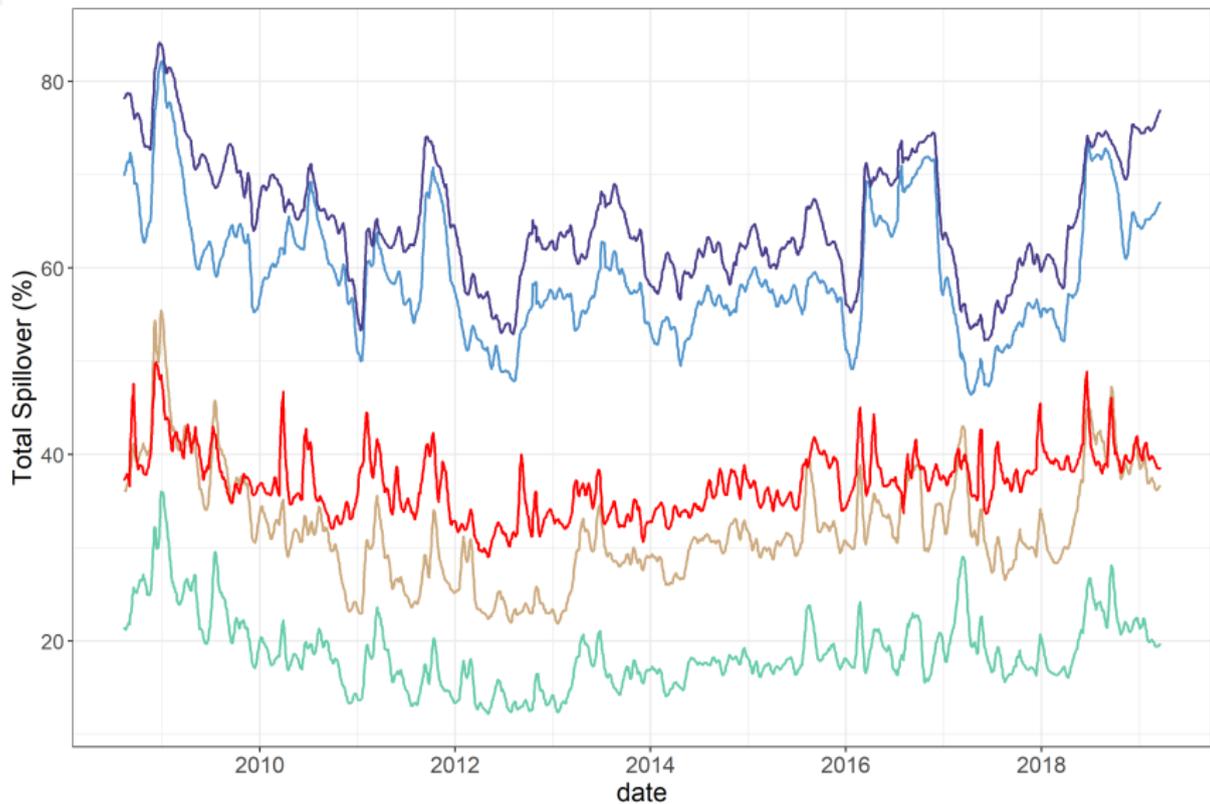
- ① VAR:  $\mathbf{y}_t = \mathbf{A}(\mathbf{L}) \mathbf{y}_{t-1} + \boldsymbol{\mu}_t$
- ② Identification:  $\mathbf{B}_0 \mathbf{y}_t = \mathbf{B}(\mathbf{L}) \mathbf{y}_{t-1} + \boldsymbol{\epsilon}_t$  and  $E_{t-1}(\boldsymbol{\epsilon}_t \boldsymbol{\epsilon}_t') = \boldsymbol{\lambda}_{|t-1}$
- ③ Use of FEVD as connectedness tables:

	$y_1$	$y_2$	$\dots$	$y_N$	To Others
$y_1$	$d_{11}^H$	$d_{12}^H$	$\dots$	$d_{1N}^H$	$\sum_{j=1}^N d_{1j}^H, j \neq 1$
$y_2$	$d_{21}^H$	$d_{22}^H$	$\dots$	$d_{2N}^H$	$\sum_{j=1}^N d_{2j}^H, j \neq 2$
$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$	$\vdots$
$y_N$	$d_{N1}^H$	$d_{N2}^H$	$\dots$	$d_{NN}^H$	$\sum_{j=1}^N d_{Nj}^H, j \neq N$
From Others	$\sum_{\substack{i=1 \\ i \neq 1}}^N d_{i1}^H$	$\sum_{\substack{i=2 \\ i \neq 2}}^N d_{i2}^H$	$\dots$	$\sum_{\substack{i=3 \\ i \neq N}}^N d_{i3}^H$	$\frac{1}{N} \sum_{\substack{i,j=1 \\ i \neq j}}^N d_{ij}^H$

# Economic Identification – FEVD

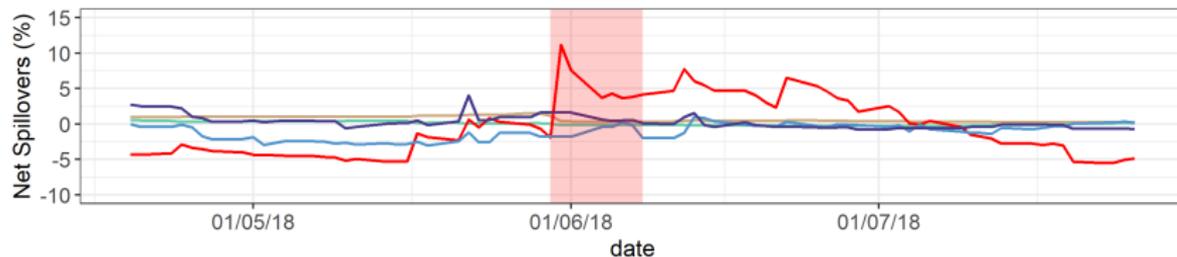
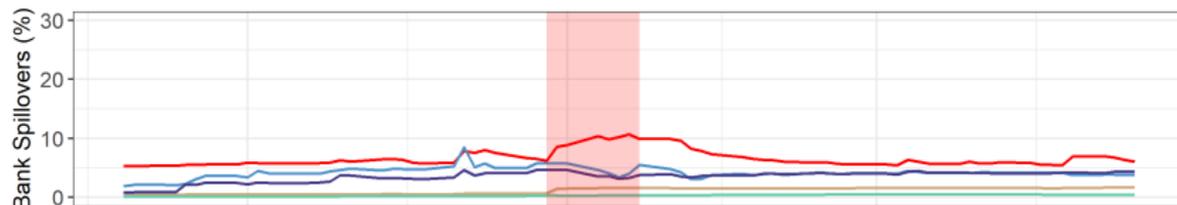
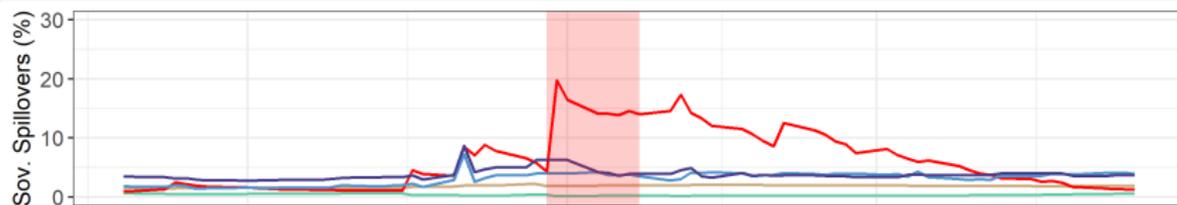
	BE_bk	FR_bk	DE_bk	IT_bk	NL_bk	ES_bk	PT_bk	DE	BE	FR	GR	NL	ES	IT	PT	IE
BE_bk	81%	1%	3%	0%	2%	0%	0%	2%	1%	1%	0%	1%	0%	0%	0%	1%
FR_bk	0%	54%	5%	2%	0%	6%	0%	2%	16%	11%	1%	2%	4%	6%	3%	5%
DE_bk	3%	9%	78%	3%	2%	3%	1%	4%	11%	7%	2%	4%	11%	8%	10%	10%
IT_bk	3%	5%	1%	83%	17%	16%	4%	0%	1%	0%	5%	0%	2%	2%	3%	0%
NL_bk	1%	2%	0%	5%	61%	1%	1%	3%	3%	3%	3%	7%	19%	4%	9%	4%
ES_bk	2%	23%	3%	1%	1%	56%	1%	10%	4%	3%	1%	1%	1%	1%	1%	4%
PT_bk	2%	1%	0%	1%	0%	1%	84%	0%	0%	1%	2%	0%	1%	1%	2%	1%
DE	4%	1%	1%	0%	2%	2%	0%	70%	13%	8%	2%	10%	4%	5%	3%	6%
BE	0%	2%	0%	0%	1%	1%	1%	1%	34%	0%	1%	3%	1%	2%	2%	5%
FR	1%	2%	0%	0%	1%	2%	1%	0%	5%	61%	1%	1%	1%	3%	1%	1%
GR	0%	0%	0%	0%	1%	0%	0%	1%	0%	0%	76%	0%	1%	0%	0%	1%
NL	1%	0%	0%	1%	1%	2%	0%	2%	9%	3%	1%	68%	1%	2%	1%	7%
ES	1%	0%	6%	2%	8%	8%	1%	1%	0%	1%	0%	1%	50%	7%	15%	1%
IT	0%	0%	0%	0%	1%	1%	0%	1%	1%	1%	2%	1%	2%	58%	3%	2%
PT	0%	0%	0%	0%	0%	0%	1%	0%	1%	0%	2%	1%	2%	1%	40%	0%
IE	1%	0%	0%	1%	0%	0%	3%	1%	0%	1%	0%	0%	1%	1%	8%	53%

# Model Comparison – Total Spillovers



— DCC Cholesky — DCC Fenger — SVAR-GARCH — VAR Cholesky — VAR GIRF

# Model Comparison – Identification (i)



— DCC Cholesky — DCC Fengler — SVAR-GARCH — VAR Cholesky — VAR GIRF

## Model Comparison – Identification (ii)

	DCC Fengler	DCC Cholesky	VAR GIRF	VAR Cholesky	<b>SVAR - GARCH</b>
(i) Subset of sovereign events	11.0	22.0	44.0	39.0	<b>78.0</b>
(ii) Total sovereign events	30.8	33.3	33.3	38.5	<b>64.1</b>
(iii) Total sovereign and bank events	34.2	36.8	39.5	47.4	<b>68.4</b>
(iv) Candelon et al. (2011)	36.3	45.4	36.4	44.6	<b>63.6</b>
(v) Alexandre et al. (2016)	12.5	25.0	75.0	0.5	<b>75.0</b>

*Note: This table reports the percentage of correct event identifications by each model.*

## Economic Validation – Theoretical channels (i)

- Do estimates also match with **theoretical contagion channels**?
- We turn to panel analysis, regressing:
  - $\bar{\omega}_{i \rightarrow j, t}$  quarterly average pairwise spillover from  $i$  to  $j$ .
  - On  $\mathbf{D}_{ijt}$  distance variables between  $i$  and  $j$ , e.g.: similarity in debt/GDP.
  - And on  $\mathbf{E}_{j \rightarrow i, t}$ , exposure variables of  $j$  on  $i$ , e.g.: trade links.
- We assess **various potential channels**, such as:
  - Trade links between sovereigns.
  - Similar portfolio channel between banks.
  - Implicit bailout channel from banks to sovereigns.
  - Balance sheet channel from sovereigns to banks...

# Economic Validation – Theoretical channels (ii)

$$\bar{\omega}_{i \rightarrow j, t} = \beta_i + D_{ijt} \beta_1 + E_{j \rightarrow i, t} \beta_2 + \alpha_t + \epsilon_{ij, t} \quad (6)$$

Regression	Expected channels	Significant channels						
<b>Sov. - Sov.</b>	Distance: Business Cycle, D/GDP Exposure: Trade, Investment	D/GDP (+), Trade (+), Investment (+)						
<b>Bank - Bank</b>	Distance: NPLs, Capital ratios Exposure: Sim. portfolios, bank claims	Capital ratios (+), Sim. portfolios (+)						
<b>Bank - Sov.</b>	<table border="0"> <tr> <td style="border-bottom: 1px solid black; padding-right: 10px;"><i>High debt</i></td> <td>Vulnerability: Capital ratio, D/GDP, Current account, GDP growth</td> <td style="border-bottom: 1px solid black; padding-right: 10px;">Cap ratio (-), D/GDP (+)</td> </tr> <tr> <td><i>Low debt</i></td> <td>Exposure: domestic sovereign / domestic NFC</td> <td>Cap ratio (-), D/GDP (-), NFC (+)</td> </tr> </table>	<i>High debt</i>	Vulnerability: Capital ratio, D/GDP, Current account, GDP growth	Cap ratio (-), D/GDP (+)	<i>Low debt</i>	Exposure: domestic sovereign / domestic NFC	Cap ratio (-), D/GDP (-), NFC (+)	
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<b>Sov. - Bank</b>	<table border="0"> <tr> <td style="border-bottom: 1px solid black; padding-right: 10px;"><i>High debt</i></td> <td>Vulnerability: NPLs, Capital ratio, Liquidity</td> <td style="border-bottom: 1px solid black; padding-right: 10px;">NPLs (+), Liquidity (-), Sovereign debt (+)</td> </tr> <tr> <td><i>Low debt</i></td> <td>Exposure: domestic sovereign / domestic NFC</td> <td>Liquidity (-), Sovereign debt (+), NFCs (+)</td> </tr> </table>	<i>High debt</i>	Vulnerability: NPLs, Capital ratio, Liquidity	NPLs (+), Liquidity (-), Sovereign debt (+)	<i>Low debt</i>	Exposure: domestic sovereign / domestic NFC	Liquidity (-), Sovereign debt (+), NFCs (+)	
<i>High debt</i>	Vulnerability: NPLs, Capital ratio, Liquidity	NPLs (+), Liquidity (-), Sovereign debt (+)						
<i>Low debt</i>	Exposure: domestic sovereign / domestic NFC	Liquidity (-), Sovereign debt (+), NFCs (+)						

# Conclusion

- Model: **Structural version** of DY based on a SVAR-GARCH with **up-to-date spillovers**.
- Methodological Results:
  - **Economic identification** enabled, even in a 16-variable system.
  - **Outperforms other models** in terms of timeliness and narrative fit.
  - Validation further supported by regression analysis: **estimates match theoretical channels**.
- Further research:
  - **Trade-off** between daily reactive estimates and time-varying  $B_0$ .
  - Fewer constraints than usual SVAR: could be used for alternative datasets (e.g. liquidity contagion).

THANK YOU

## BACKUP SLIDES

## Annex – FEVD build-up

- With the MA representation of  $\mathbf{Y}_t$ :

$$\mathbf{Y}_{t+H} - \mathbf{Y}_{t+H|t} = \sum_{i=0}^{H-1} \Theta_i \epsilon_{t+H-i} \quad (7)$$

- FEVD is a function of MSPE:

$$\begin{aligned} MSPE_t(H) &= E_t(\mathbf{Y}_{t+H} - \mathbf{Y}_{t+H|t})(\mathbf{Y}_{t+H} - \mathbf{y}_{t+H|t})' \\ &= \sum_{i=0}^{H-1} \Theta_i \lambda_{t+H-i|t}^* \Theta_i' \end{aligned} \quad (8)$$

- So connectedness may come from 2 sources, propagation mechanisms ( $\Theta_i = \mathbf{J} \mathbf{A}^i \mathbf{J}' \mathbf{B}_0^{-1}$ ) or volatility ( $\lambda_{t+H-i|t}^*$ ). Volatility being projected at each  $t$  with the GARCH structure of the SVAR.

## Annex – Theoretical channels, between sovereigns

$$\bar{\omega}_{i \rightarrow j, t}(h) = \beta_i + \alpha_t + \beta_2 d_{ij, t}^{GDP} + \beta_3 d_{ij, t}^{\frac{D}{GDP}} + \beta_4 \text{exposure}_{j \rightarrow i, t}^k + \epsilon_{ij, t} \quad (9)$$

	(1)	(2)	(3)
Similar BC	-0.00005 (0.001)	-0.002 (0.001)	-0.001 (0.001)
Similar D/GDP	0.021*** (0.002)	0.007*** (0.002)	0.014*** (0.002)
Trade exposure		0.433*** (0.020)	
Investment exposure			0.236*** (0.028)
Time fixed effects?	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>
i fixed effects?	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>
Observations	3,240	3,240	3,171
R <sup>2</sup>	0.448	0.598	0.490
Adjusted R <sup>2</sup>	0.438	0.591	0.481

\*p<0.05; \*\*p<0.01; \*\*\*p<0.001

# Annex – Theoretical channels, between banks

$$\bar{\omega}_{i \rightarrow j, t}(h) = \beta_i + \alpha_t + \beta_2 d_{ij, t}^{NPL} + \beta_3 d_{ij, t}^{Lev.R.} + \beta_4 \text{exposure}_{j \rightarrow i, t}^k + \epsilon_{ij, t} \quad (10)$$

	(1)	(2)	(3)
Similar NPLs	-0.001 (0.003)	-0.003 (0.003)	-0.001 (0.003)
Sim. Capital ratios	0.037*** (0.009)	0.029** (0.010)	0.040*** (0.010)
Similar portfolio		7.304*** (1.355)	
Bank claims			-0.013 (0.009)
Time fixed effects?	Yes	Yes	Yes
i fixed effects?	Yes	Yes	Yes
Observations	1,812	1,812	1,812
R <sup>2</sup>	0.434	0.439	0.435
Adjusted R <sup>2</sup>	0.417	0.422	0.417

\* p<0.05; \*\* p<0.01; \*\*\* p<0.001

# Annex – Theoretical channels, bank to sovereign

$$\bar{\omega}_{bank_i \rightarrow sov_i, t}(h) = \beta_0 + \alpha_t + \beta_1 v_{bank_i, t}^{Lev.R.} + \beta_2 v_{sov_i, t}^{D/GDP} + \beta_3 v_{sov_i, t}^{CA} + \beta_4 v_{sov_i, t}^{gGDP} + \beta_5 \text{exposure}_{sov_i \rightarrow bank_i, t}^k + \epsilon_{bank_i, sov_i, t} \quad (11)$$

	<i>High debt countries</i>			<i>Low debt countries</i>		
	(1)	(2)	(3)	(4)	(5)	(6)
Capital	-0.91*** (0.21)	-0.91*** (0.21)	-1.01** (0.31)	-5.01*** (1.04)	-5.07*** (1.05)	-3.88*** (1.07)
Debt to GDP	1.00*** (0.17)	1.01*** (0.17)	0.88*** (0.21)	3.74** (1.29)	3.56** (1.28)	3.50** (1.22)
Current Account	0.13 (0.20)	0.13 (0.26)	0.23 (0.28)	-0.63 (1.26)	-0.62 (1.28)	0.16 (1.31)
GDP growth	-0.06 (0.22)	-0.05 (0.23)	-0.05 (0.22)	-0.29 (0.66)	-0.31 (0.66)	-0.28 (0.68)
Sov. exposure		0.06 (1.92)			2.94 (9.63)	
Non-bank exposure			-0.88 (1.50)			20.52*** (6.02)
Time fixed effects?	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>
Observations	174	174	174	123	123	123
R <sup>2</sup>	0.74	0.74	0.74	0.93	0.93	0.94
Adjusted R <sup>2</sup>	0.64	0.64	0.64	0.89	0.89	0.90

\*p<0.05; \*\*p<0.01; \*\*\*p<0.001

# Annex – Theoretical channels, sovereign to bank

$$\bar{\omega}_{sov_i \rightarrow bank_i, t}(h) = \beta_0 + \alpha_t + \beta_1 v_{bank_i, t}^{NPL} + \beta_2 v_{bank_i, t}^{Lev.R.} + \beta_3 v_{bank_i, t}^{Liq.R.} + \beta_5 \text{exposure}_{bank_i \rightarrow sov_i, t}^k + \epsilon_{sov_i, bank_i, t} \quad (12)$$

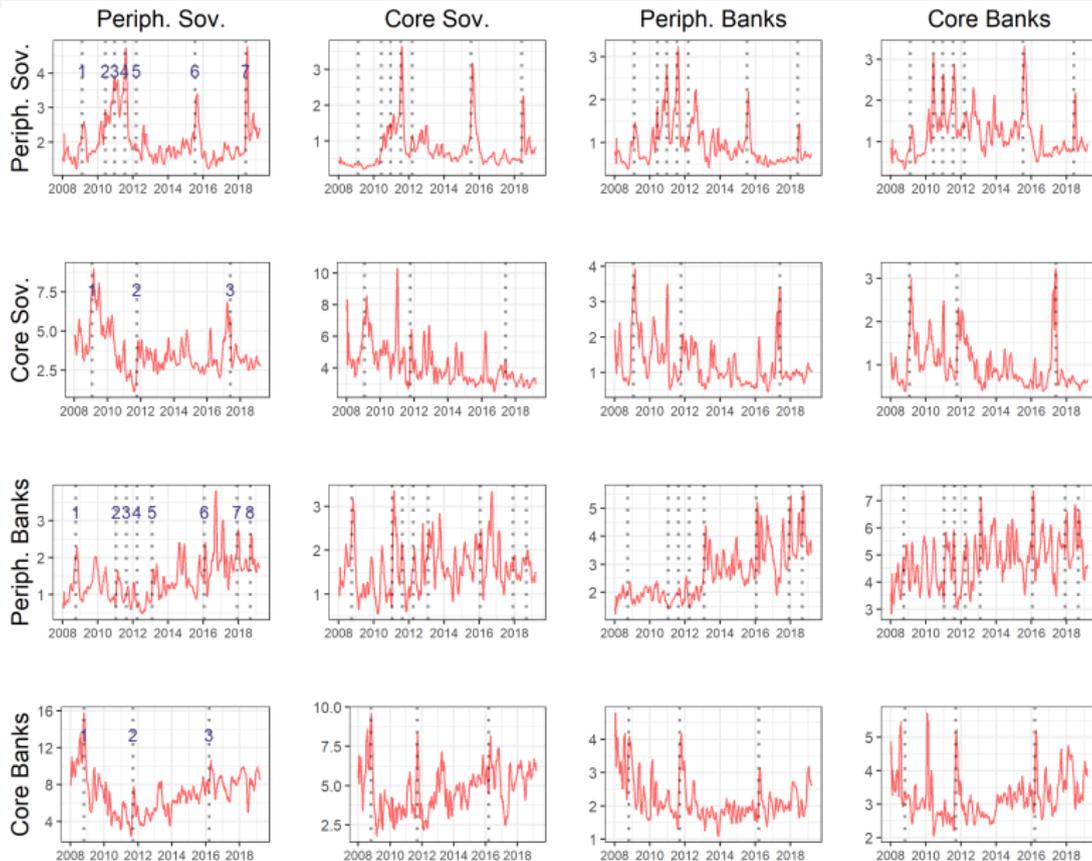
	<i>High debt countries</i>			<i>Low debt countries</i>		
	(1)	(2)	(3)	(4)	(5)	(6)
NPLs	2.61** (0.81)	2.24** (0.81)	2.74** (0.84)	-0.10 (0.18)	-0.12 (0.17)	-0.19 (0.17)
Capital	-1.38 (0.77)	-0.80 (0.76)	-2.49 (1.31)	-0.06 (0.23)	0.34 (0.23)	0.41 (0.31)
Liquid assets	-4.03*** (0.48)	-6.17*** (1.01)	-4.60*** (0.69)	-0.92*** (0.18)	-0.11 (0.29)	-0.43 (0.29)
Exposure domestic gov. debt		0.18** (0.07)			0.18*** (0.05)	
Exposure domestic NFCs			-0.04 (0.03)			0.05** (0.02)
Time fixed effects?	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>	<i>Yes</i>
Observations	174	174	174	121	121	121
R <sup>2</sup>	0.54	0.55	0.54	0.88	0.89	0.88
Adjusted R <sup>2</sup>	0.37	0.38	0.37	0.80	0.82	0.81

\*p<0.05; \*\*p<0.01; \*\*\*p<0.001

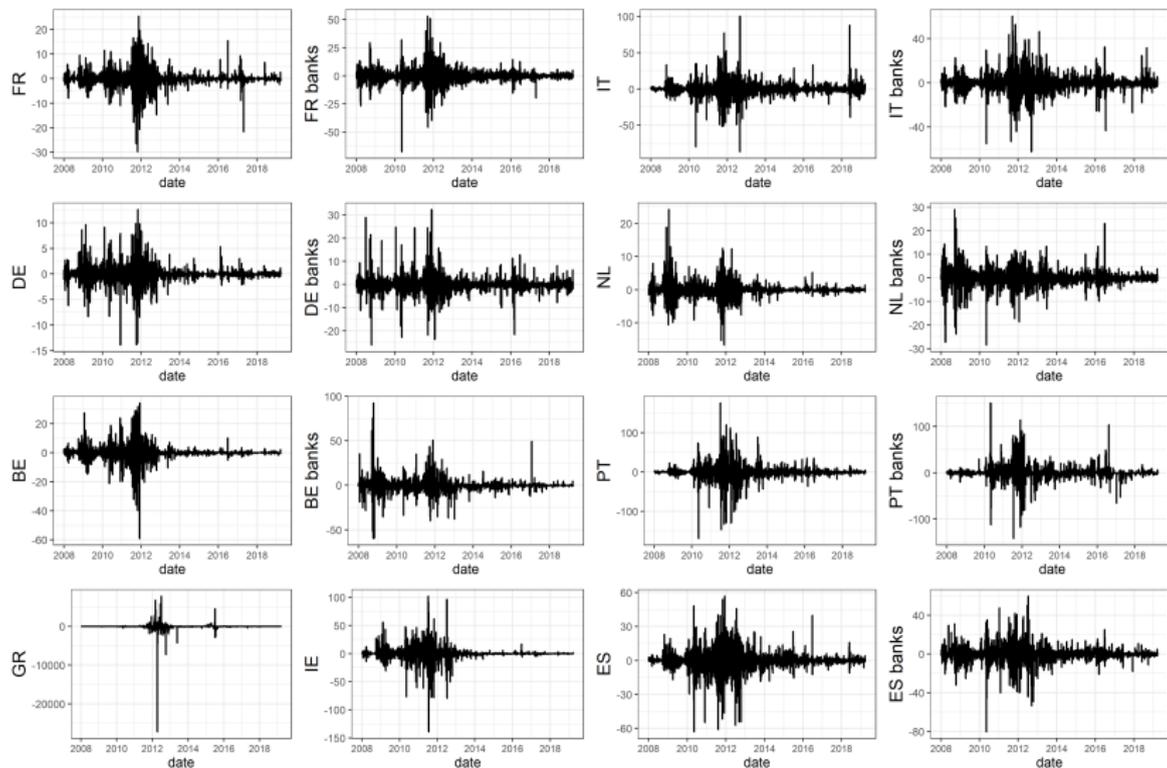
## Annex – Granger causality

$H_0$ : SVAR-GARCH does not Granger cause:	F - test	p-value
<i>Rolling window estimated models</i>		
VAR Cholesky	14.91	3.627e-07***
VAR GIRF	6.0492	0.002391 **
<i>GARCH-related models</i>		
DCC Cholesky	1.0527	0.3491
DCC Fengler	1.3641	0.2558
$H_0$ : SVAR-GARCH is not Granger caused by:	F - test	p-value
<i>Rolling window estimated models</i>		
VAR Cholesky	0.5159	0.597
VAR GIRF	0.9483	0.3875
<i>GARCH-related models</i>		
DCC Cholesky	0.4206	0.6567
DCC Fengler	8.8071	0.0001539***

# Annex – Pairwise spillovers



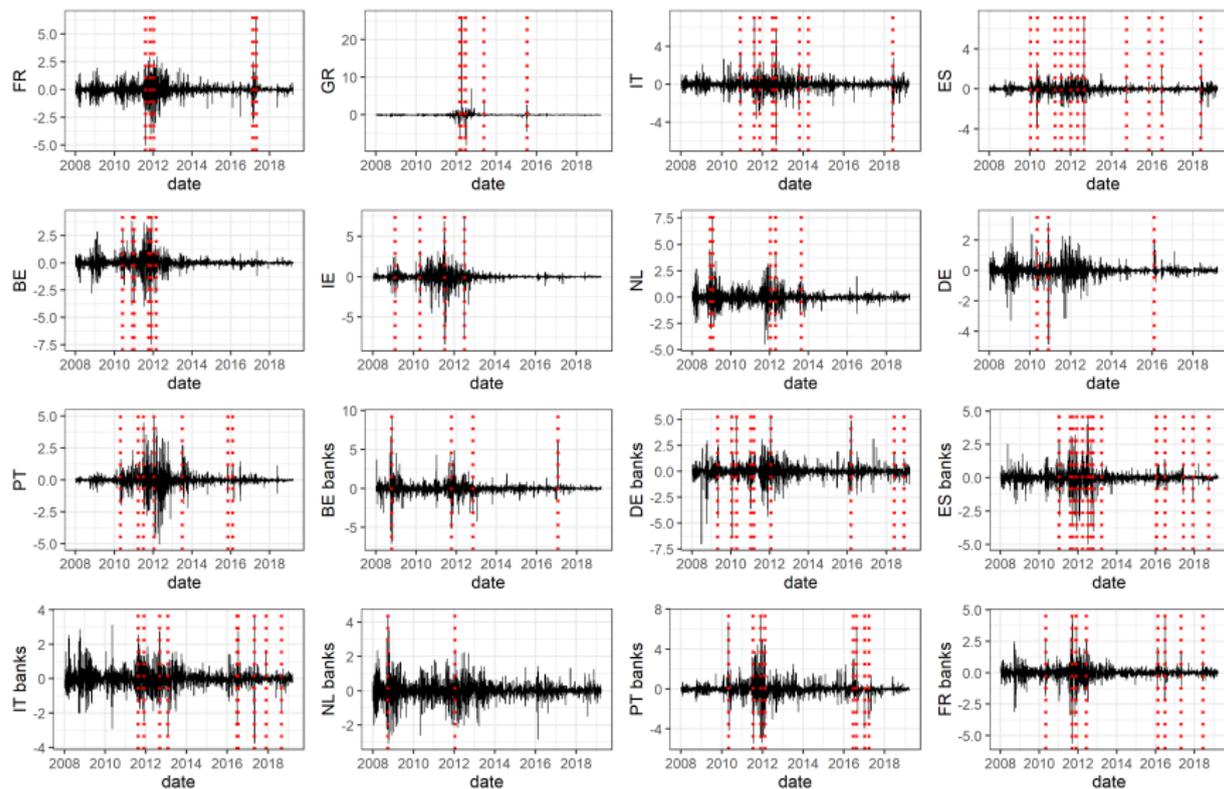
# Annex – Inputs



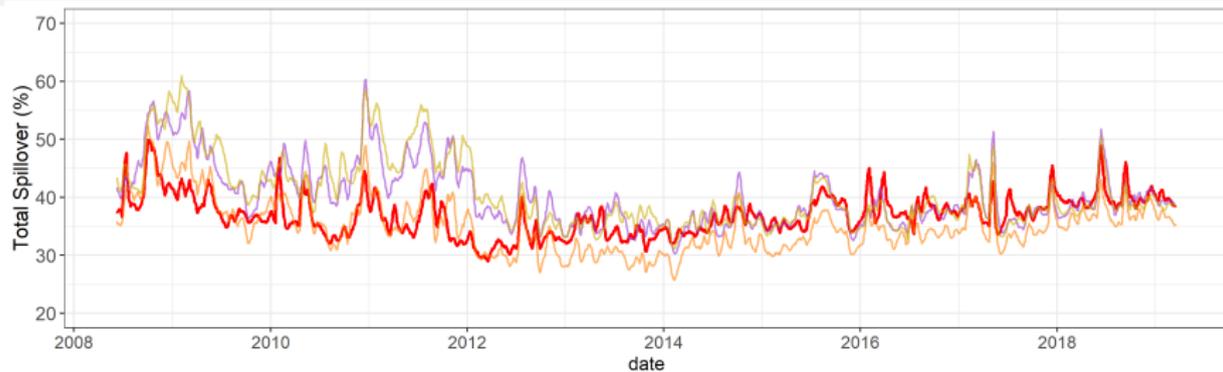
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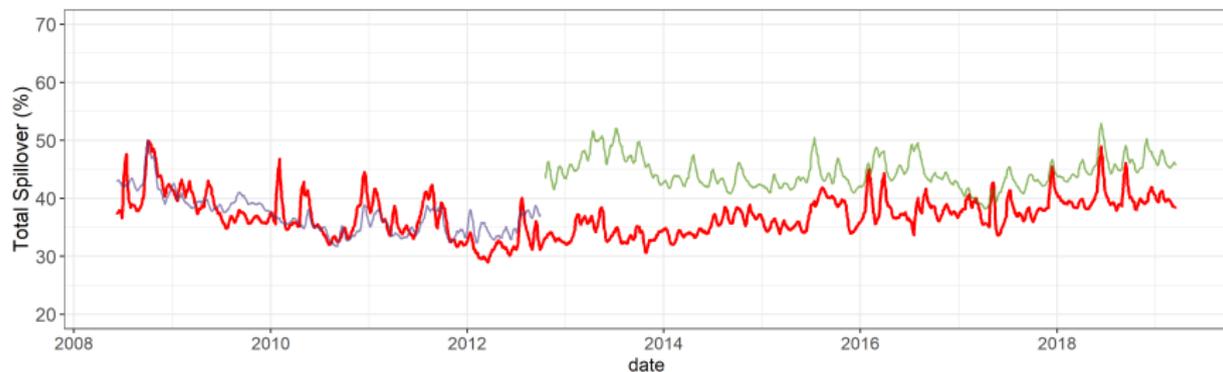
# Annex – Events



# Annex – Robustness



— Citi-Oil-US & UK sov. — Citi-Oil-US & UK sov.-US & UK banks — Citi-Oil-US sov.-US banks — SVAR-GARCH



— Fratzscher et al. 1 — Fratzscher et al. 2 — SVAR-GARCH

## Annex – Test identification

$h$ under $H_0$	$Q_1(1)$	df	p-value
1	124.3405	1	$< 10^{-5}$
2	113.4685	1	$< 10^{-5}$
3	85.0733	1	$< 10^{-5}$
4	66.6269	1	$< 10^{-5}$
5	60.7231	1	$< 10^{-5}$
6	46.2298	1	$< 10^{-5}$
7	38.0658	1	$< 10^{-5}$
8	35.8007	1	$< 10^{-5}$
9	25.3033	1	$< 10^{-5}$
10	16.2284	1	5.615e-05
11	13.3168	1	0.000263
12	12.6034	1	0.000385
13	517.7083	1	$< 10^{-5}$
14	185.0355	1	$< 10^{-5}$
15	154.8558	1	$< 10^{-5}$